

Binomial Tails Domination for Random Graphs via Bell Polynomials

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ABSTRACT We obtain an inequality between tails of binomial distributions by establishing a combinatorial identity for a sum involving binomial and multinomial coefficients associated with the Bell polynomials. As an application a connection to random graphs is presented.

Keywords Bell polynomials, Di Bruno formula, Random graphs.

1. Introduction

As remarked in [5] the task of reducing combinatorial sums to a single closed form expression is an art in itself. Indeed, despite various algorithms and techniques developed over the years, quite often one encounters a situation in which the known methodology does not apply and requires a novel approach. This paper is an example of such a case in which a combinatorial identity emerged as a by-product of studying polynomials associated with binomial random graphs. Namely, let $Q(x)$ be a binomial-type polynomial defined by

$$Q(x) = \sum_{i=0}^m \binom{n}{i} (x^k - 1)^i, \quad 0 \leq m \leq n, \quad k \geq 1 \quad (1)$$

and consider

$$Q^j(1) = \left. \frac{d^j Q(x)}{dx^j} \right|_{x=1}, \quad 0 \leq j \leq m. \quad (2)$$

Evaluating $Q^j(1)$ when $k = 1$ is trivial however for $k \geq 2$ the matter is far less obvious.

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Consequently it is of interest to find an explicit representation for $Q^{(j)}(1)$ and explore the possibility of reducing it to a simple closed form. It turns out that the first part can be achieved by applying *di Bruno formula* for differentiation of composite function while the second part amounts to showing that so obtained representation can in fact be simplified to a single expression referred to as the combinatorial identity for *Bell polynomials*.

2. The Bell Polynomials Identity

We begin by recalling some facts concerning the *Bell polynomials* [6]. Let $B_j(x_1, \dots, x_j)$ be the j -th exponential *Bell polynomial* defined by the generating function

$$e^{\sum_{i=1}^{\infty} x_i \frac{t^i}{i!}} = 1 + \sum_{j=1}^{\infty} B_j(x_1, \dots, x_j) \frac{t^j}{j!}$$

where

$$\begin{aligned} B_j(x_1, \dots, x_j) &= \sum_{l_1+2l_2+\dots+jl_j=j} \frac{j!}{(1!)^{l_1} l_1! \dots (j!)^{l_j} l_j!} x_1^{l_1} \dots x_j^{l_j} \\ &= \sum_{l=1}^j \left(\sum_{l_1+2l_2+\dots+jl_j=l} \frac{j!}{(1!)^{l_1} l_1! \dots (j!)^{l_j} l_j!} x_1^{l_1} \dots x_j^{l_j} \right) \\ &\equiv \sum_{l=1}^j B_{j,l}(x_1, \dots, x_j) \end{aligned} \quad (3)$$

and $\{l_1, l_2, \dots, l_j\}$ are all nonnegative integer solutions of the *Diophantine equation* $l_1 + 2l_2 + \dots + jl_j = j$. Below are the first few polynomials

$$B_1(x_1) = x_1$$

$$B_2(x_1, x_2) = x_2 + x_1^2$$

$$B_3(x_1, x_2, x_3) = x_3 + 3x_2x_1 + x_1^3$$

$$B_4(x_1, x_2, x_3, x_4) = x_4 + (4x_3x_1 + 3x_2^2) + 6x_2x_1^2 + x_1^4$$

$$B_5(x_1, x_2, x_3, x_4, x_5) = x_5 + (5x_4x_1 + 10x_3x_2) + (10x_3x_1^2 + 15x_2x_1) + 10x_2^2x_1^3 + x_1^5$$

where the right hand side is written as a sum $B_{j,1}(x_1, \dots, x_j) + \dots + B_{j,j}(x_1, \dots, x_j)$, with parenthesis used as needed, of the associated partial *Bell polynomials* $\{B_{j,l}(x_1, \dots, x_j), l = 1, \dots, j\}$. Notice that for $x_1 = \dots = x_j \equiv x$, $B_j(x_1, \dots, x_j)$ turns into the classical *Bell polynomial* and when $x \equiv 1$ then $B_{j,l}(1, \dots, 1) = S(j, l)$, $B_j(1, \dots, 1) = \sum_{l=1}^j S(j, l) = B_j$ which are the *Stirling numbers* of the second kind and the *Bell numbers* respectively corresponding to all possible partitions of a j -element set into l blocks of its non-empty disjoint subsets, $l = 1, \dots, j$. Furthermore, *di Bruno formula* for the j -th derivative of the composite function $f(g(x))$ reads

$$f(g(x))^{(j)} = \sum_{l_1+2l_2+\dots+jl_j=j} \frac{j!}{(1!)^{l_1} l_1! \dots (j!)^{l_j} l_j!} (g^{(1)})^{l_1} \dots (g^{(j)})^{l_j} f^{(l_1+l_2+\dots+l_j)}$$

$$\begin{aligned}
 &= \sum_{l=1}^j \left(\sum_{\substack{l_1+l_2+\dots+l_j=l \\ l_1+2l_2+\dots+jl_j=j}} \frac{j!}{(1!)^{l_1} l_1! \dots (j!)^{l_j} l_j!} (g^{(1)})^{l_1} \dots (g^{(j)})^{l_j} \right) f^{(l)} \quad (4) \\
 &\equiv \sum_{l=1}^j f^{(l)} B_{j,l}(g^{(1)}, \dots, g^{(j)})
 \end{aligned}$$

where the last equality comes from (3) while $f^{(l)} = \frac{d^l f(y)}{dy^l} |_{y=g(x)}$ and $g^{(l)} = \frac{d^l g(x)}{dx^l}$, $l = 1, 2, \dots, j$.

Lemma 1 *Let $(k)_l = k(k - 1)\dots(k - l + 1)$ denote the falling factorial. Then for any positive integers i, j, k we have*

$$((x^k - 1)^i)^{(j)}(1) = \begin{cases} i! B_{j,i}((k)_1, \dots, (k)_j), & j \geq i \\ 0, & j < i \end{cases} \quad (5)$$

Proof. Applying (4) for $f(y) = y^i$ and $y = g(x) = x^k - 1$ gives

$$((x^k - 1)^i)^{(j)} = \sum_{l=1}^j f^{(l)} B_{j,l}(g^{(1)}, \dots, g^{(j)}) \quad (6)$$

where $g^{(l)} = (k)_l x^{k-l}$ for $l \leq k$ and $g^{(l)} = 0 = (k)_l$ for $l > k$. Since

$$f^{(l)} = ((x^k - 1)^i)^{(l)} = \begin{cases} (i)_l (x^k - 1)^{i-l}, & l = 1, \dots, j, & j < i \\ (i)_l (x^k - 1)^{i-l}, & l = 1, \dots, i, & j \geq i \\ 0, & l = i + 1, \dots, j, & j \geq i \end{cases}$$

therefore

$$f^{(l)}(1) = \begin{cases} 0, & l = 1, \dots, j, & j < i \\ i!, & l = i, & l \leq j, & j \geq i \\ 0, & l \neq i, & l \leq j, & j \geq i \end{cases}$$

which combined with (6) yields

$$((x^k - 1)^i)^{(j)}(1) = \begin{cases} i! B_{j,i}((k)_1, \dots, (k)_j), & j \geq i \\ 0, & j < i \end{cases}$$

and the proof is complete.

Lemma 2 *Let $1 \leq j \leq m \leq n$, $k \geq 1$ and $Q(x)$ and $Q^{(j)}(1)$ be as defined by (1)-(2). Then*

$$Q^{(j)}(1) = (nk)_j \quad (7)$$

Proof. Since

$$x^{nk} = (1 + x^k - 1)^n = \sum_{i=0}^m \binom{n}{i} (x^k - 1)^i + \sum_{i=m+1}^n \binom{n}{i} (x^k - 1)^i$$

it follows

$$Q(x) = x^{nk} - \sum_{i=m+1}^n \binom{n}{i} (x^k - 1)^i = x^{nk} - (x - 1)^{m+1} P(x)$$

whence Leibnitz's differentiation $\{(x - 1)^{m+1} P(x)\}^{(j)}(1) = 0$ for $1 \leq j \leq m$ gives

$$Q^{(j)}(1) = (nk)_j$$

as desired.

Theorem 1 (Bell polynomials identity). *For any positive integers i, j, k the following combinatorial identity holds true*

$$\sum_{i=1}^j \binom{n}{i} \sum_{\substack{i_1+i_2+\dots+i_j=i \\ i_1+2i_2+\dots+j i_j=j}} \binom{i}{i_1, i_2, \dots, i_j} \binom{k}{1}^{i_1} \dots \binom{k}{j}^{i_j} = \binom{nk}{j} \quad (8)$$

or equivalently upon multiplying (8) by $j!$

$$\sum_{i=1}^j (n)_i B_{j,i}((k)_1, \dots, (k)_j) = (nk)_j \quad (9)$$

where $B_{j,i}(x_1, \dots, x_j) = \sum_{\substack{i_1+i_2+\dots+i_j=i \\ i_1+2i_2+\dots+j i_j=j}} \frac{j!}{(1!)^{i_1} i_1! \dots (j!)^{i_j} i_j!} x_1^{i_1} \dots x_j^{i_j}$ are the partial Bell polynomials defined in (3).

Proof. It suffices to verify (9) but it is a straightforward consequence of Lemma 1 and Lemma 2 because for $1 \leq j \leq m \leq n$ we have

$$\begin{aligned} (nk)_j &= Q^{(j)}(1) = \sum_{i=0}^m \binom{n}{i} ((x^k - 1)^i)^{(j)}(1) = \sum_{i=0}^j \binom{n}{i} ((x^k - 1)^i)^{(j)}(1) \\ &= \sum_{i=1}^j (n)_i B_{j,i}((k)_1, \dots, (k)_j) \end{aligned}$$

This concludes the proof of identities (8)-(9).

3. Domination of Binomial Tails

In the sequel we shall focus on comparison of random variables X and Y in the following [7] sense.

Definition We say that X is stochastically larger than Y (X dominates Y), written $X \succ Y$, if the tail probability of X dominates the tail probability of Y according to

$$P(X > a) \geq P(Y > a) \text{ for all } a, \text{ written } F \succ G \quad (10)$$

where F and G are the distributions of X and Y respectively.

Notice that for nonnegative integer-valued random variables, thanks to $P(X \geq 0) = P(Y \geq 0) = 1$, (10) is equivalent to $P(X \geq k) \geq P(Y \geq k)$, $k = 1, 2, \dots$ and implies $EX = \sum_{k \geq 1} P(X \geq k) \geq \sum_{k \geq 1} P(Y \geq k) = EY$ for the expected values.

In what follows $X \sim b(m, \theta)$ stands for *binomial distribution* with parameter θ , i.e., $P(X = i) = \binom{m}{i} \theta^i (1 - \theta)^{m-i}$, $i = 0, 1, \dots, m$.

Theorem 2 Let $X \sim b(nk, p)$ and $Y \sim b(n, 1 - (1 - p)^k)$ then $X \succ Y$ or equivalently

$$b(nk, p) \succ b(n, 1 - (1 - p)^k) \quad (11)$$

for any positive integers n, k and $0 \leq p \leq 1$.

Proof. Since $p = 1$ is trivial, it suffices to show that for $0 \leq p < 1$, $0 \leq m \leq n$, $k \geq 1$

$$\begin{aligned} P(X \geq m + 1) &= \sum_{i=m+1}^{nk} \binom{nk}{i} p^i (1 - p)^{nk-i} \geq P(Y \geq m + 1) \\ &= \sum_{i=m+1}^n \binom{n}{i} ((1 - (1 - p)^k)^i ((1 - p)^k)^{n-i} \end{aligned}$$

or equivalently, by $P(X \leq m) = 1 - P(X \geq m + 1) \leq 1 - P(Y \geq m + 1) = P(Y \leq m)$, that

$$\sum_{i=0}^m \binom{n}{i} ((1 - (1 - p)^k)^i ((1 - p)^k)^{n-i} \geq \sum_{i=0}^m \binom{nk}{i} p^i (1 - p)^{nk-i} \quad (12)$$

Dividing (12) by $(1 - p)^{nk}$, setting $x = \frac{1}{1-p} \geq 1$, and using (1) gives

$$Q(x) = \sum_{i=0}^m \binom{n}{i} (x^k - 1)^i \geq \sum_{i=0}^m \binom{nk}{i} (x - 1)^i, \quad x \geq 1 \quad (13)$$

To verify (13), following the proof of Lemma 2, we express $Q(x)$ in two different ways

$$Q(x) = x^{nk} - \sum_{i=m+1}^n \binom{n}{i} (x^k - 1)^i = \sum_{j=0}^m a_j (x - 1)^j + \sum_{j=m+1}^{mk} a_j (x - 1)^j \quad (14)$$

where the last equality is the representation of the mk -th degree polynomial $Q(x)$ in the basis $\{(x - 1)^j, j = 0, 1, \dots, mk\}$. Now by *Taylor's* expansion combined with (2) and (7) of Lemma 2 we have

$$a_j = \frac{Q^{(j)}(1)}{j!}, \quad j = 0, 1, \dots, mk \text{ and } a_j = \binom{nk}{j}, \quad j = 0, 1, \dots, m \quad (15)$$

whence (1), (14)-(15) gives

$$\sum_{i=0}^m \binom{n}{i} (x^k - 1)^i = \sum_{i=0}^m \binom{nk}{i} (x - 1)^i + \sum_{j=m+1}^{mk} a_j (x - 1)^j, \quad x \geq 1 \quad (16)$$

which proves (13) because the second sum in the right hand side of (16) is nonnegative thanks to

$$a_j = \frac{Q^{(j)}(1)}{j!} = \frac{(\sum_{i=0}^m \binom{n}{i} (x^k - 1)^i)^{(j)}(1)}{j!} \geq 0, \quad j = m + 1, \dots, mk$$

which follows from the fact that $((x^k - 1)^i)^{(j)}(1) \geq 0$ by (5) of Lemma 1, since $B_{j,i}(x_1, \dots, x_j)$ as a polynomial with positive integer coefficients satisfies $B_{j,i}(x_1, \dots, x_j) \geq 0$ whenever $x_1, \dots, x_j \geq 0$. □

Before we present an application of (10)-(11) domination, let us remark that it extends to the case where G is a conditional distribution of Y given a random vector, say Z , since $F \succ G$ is the distribution property.

4. Application to Random Graphs

For a reference on the subject of random graphs the reader may consult a monograph [1] while some recent findings concerning binomial random graphs and their representations can be found in [3], [4]. A *binomial* random graph $G[V, p]$ on the set of vertices $V = \{v_1, \dots, v_N\}$ is built by placing a connecting edge between any pair of $\binom{N}{2}$ vertices independently at random with probability p and placing no edge with probability $1 - p$. Let $Z = \#$ edges in $G[V, p]$, then Z has the *binomial* distribution $b\left(\binom{N}{2}, p\right)$ according to $P(Z = i) = \binom{\binom{N}{2}}{i} p^i (1 - p)^{\binom{N}{2} - i}$, $i = 0, 1, \dots, \binom{N}{2}$.

Let V_1 and V_2 be two disjoint non-empty subsets of V such that $|V_1| = n, |V_2| = k, n + k \leq N$. A *bipartite* random graph $G[V, V_1, V_2, p]$ is built by placing connecting edges between nk pairs of vertices from V_1 and V_2 with probability p . We define two random variables associated with $G[V, V_1, V_2, p]$ First of which.

$$X = \# \text{ of edges between } V_1 \text{ and } V_2 \quad (17)$$

has the *binomial* distribution $b(nk, p)$ whereas the second is defined as follows. For any vertex $v \in V_1$ $P(\text{there is a connecting edge from } v \text{ to at least one vertex from } V_2) = 1 - (1 - p)^k$. Then

$$Y = \# \text{ of vertices in } V_1 \text{ that connect to at least one vertex from } V_2 \quad (18)$$

has the binomial distribution $b(n, 1 - (1 - p)^k)$.

Theorem 3 Consider $G[V, p]$ on the set of vertices $V = \{v_1, \dots, v_n\}$. Given $x \in V$, let $\Gamma_k(x) = \{y \in V \mid d(x, y) = k\}$, where $d(x, y)$ is the distance between the vertices $\{x, y\}$ (= length of the shortest chain of edges connecting x and y) and let $N_k(x) = \cup_{i=0}^k \Gamma_i(x) \subsetneq V$. Let

$Z = (|\Gamma_k(x)|, |N_k(x)|)$, $Y = |\Gamma_{k+1}(x)|$ and $G(a) = G_{Y|Z}(a|Z) = P(Y \leq a|Z)$ be the conditional distribution. Then we have the domination

$$\begin{aligned} b(|\Gamma_k(x)|(n - |N_k(x)|), p) &> G_{|\Gamma_{k+1}(x)| | (|\Gamma_k(x)|, |N_k(x)|)} \\ &= b\left(n - |N_k(x)|, 1 - (1 - p)^{|\Gamma_k(x)|}\right) \end{aligned} \quad (19)$$

for $0 \leq k < n$, $0 \leq p \leq 1$.

Proof. Apply Theorem 2 with $|V_1| = n - |N_k(x)|$, $|V_2| = |\Gamma_k(x)|$ and observe that the distribution of $|\Gamma_{k+1}(x)|$, given $(|\Gamma_k(x)|, |N_k(x)|)$, corresponds to the bipartite random graph $G[V, V \setminus N_k(x), \Gamma_k(x), p]$. \square

A special case of domination (19) under specific restrictions on the range of $p = p(n)$ was a basis for several results in [2]. Our theorem shows that (19) holds true universally without any restrictions on $0 \leq p \leq 1$ whatsoever and may pave the way toward new probability estimates for the upper bound of $|\Gamma_k(x)|$ for the associated random graphs.

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